



Thailand Statistician
July 2021; 19(3): 583-592
<http://statassoc.or.th>
Contributed paper

A Comparative Study of Bayesian and Frequentist Testing for Model Comparison in Nested Model

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Received: 6 May 2020

Revised: 16 June 2020

Accepted: 11 August 2020

Abstract

Comparison of the two modelling approaches; frequentist and Bayesian had been carried out by various researchers in the area of estimation of parameters in regression model. Modelers are also interested in comparing different models to know the correct one. However, the use of relevant prior information about the data in the concept of model comparison for Bayesian can help to select the right model. This work facilitates comparison between frequentist and Bayesian methods in a nested model when all the parameters considered are evaluated at zero through a simulation study. The results from both simulation and real life data suggest that Bayesian approach with the use of Savage-Dickey density ratio (SDDR) provide a reasonable decision for the nested model than frequentist approach. It was also demonstrated that the SDDR is a true representation of Bayes factor.

Keywords: Bayes factor, Bayesian, frequentist, prior, Savage-Dickey density ratio.

1. Introduction

In statistical modelling, apart from estimation of parameters of models, some of the things a model builder would also wish to do are to compare different models. However, choosing the correct model by model builders and researchers can be difficult. Model comparison is a way by which models are compared to one another and the best model is chosen.

There are so many works that had been carried out on model comparison in Classical point of views. According to Harvey (1990), some classical model comparisons methods are Akaike's Information Criterion (AIC), Mallows's C_p criterion, Amemiya's Prediction criterion (APC), Davidson-Mackinon test etc. Most of these aforementioned classical methods of model comparison are used purposely for non-nested model.

In Bayesian framework, there are limited applications of model comparison and also Bayesian inference is more recent than classical method. Thus, some prominent works on Bayesian method of model comparison are; Smith and Spiegelhalter (1980), Griffiths and Wan (1994), O'Hagan (1995), Kass and Wasserman (1995), Dickey et al. (2003), Plummer (2008), Feng and Giles (2009), Wetzels et al. (2010), and Li et al. (2017).

Smith and Spiegelhalter (1980) looked into forms of model choice criteria and compared alternative nested linear models on the basis of their asymptotic properties. Bayesian estimation of normal regression model with an uncertain inequality constraint was considered by Griffiths and Wan (1994). They adopted a non-informative prior and uncertainty concerning the inequality restriction that was represented by prior odds ratio while O'Hagan (1995) proposed a fractional Bayes factor for Bayesian comparison of models. The approach was found to be consistent, simple, robust and coherent.

Carlin and Chib (1995) summarizes a range of computational methods for obtaining estimates of Bayes factor. They provide a convenient method by using a marginal likelihood for comparing models by their fit with less theoretical problems attached to it than encountered when comparing non-nested models in a classical framework.

A generalized method Savage-Dickey Density Ratio (SDDR) for computing a Bayes factor was developed by Verdinelli and Wasserman (1995). It was observed that their methods in terms of computational complexity can be extended to other models. Kass and Wasserman (1995) computed a Bayes factor for testing equality restriction in the presence of nuisance parameter priors. Their results suggested that Schwartz criterion can provide sensible approximate solutions to Bayesian testing problems when the hypotheses are nested.

Several methods of estimating Bayes factors for non-linear models when it is possible to simulate observations from posterior distributions through MCMC and other techniques was examined by Diciccio et al. (1997). The simulated versions of Laplace's, Bartlett correction, importance sampling and reciprocal importance sampling techniques were considered. They found out that simulated version of Laplace's method was the most accurate approach among all the techniques.

A deviance-based loss function was derived by Plummer (2008) using a decision-theoretic framework. This approach was developed in order to capture some mixture models. The theoretical properties of the loss functions were examined in normal linear models and exponential family models using this penalized loss function. This approach was applied in mixture modelling and disease mapping.

Feng and Giles (2009) employed a posterior odds analysis to select the correct number of clusters for a Bayesian fuzzy regression analysis. They used a natural conjugate prior for parameters and concluded that the Bayesian posterior odds can provide a very powerful tool for choosing the number of clusters through their results.

Wetzel et al. (2010) in their work proposed an Encompassing Prior (EP) approach to facilitate Bayesian model selection for nested models with inequality constraints. EP approach generalizes the Savage-Dickey ratio method and can accommodate both inequality and exact equality constraints. Their EP approach was found to be computationally efficient procedure for calculating Bayes factor for nested models.

The plug-in predictive distribution as an alternative to DIC was also proposed by Li et al. (2017). This method was purposely to provide an asymptotically unbiased estimation to the new expected Kullback-Leibler (KL) divergence under a general framework. It was found out that this alternative DIC is easy for computation from Markov Chain Monte Carlo (MCMC) output and also has a smaller penalty term than the original DIC.

Most authors had considered the estimation of parameters of regression model by facilitating comparison between classical and Bayesian approaches (See Zellner 1976, Kleibergen and Zivot 2003, Adepoju and Ojo 2018 etc.). However, researchers are also interested in knowing the best method for choosing among competing models or hypotheses of a given phenomenon.

A model is said to be nested when one model is a special case of another model. Regression in a nested model for model comparison is of two types. The first entails comparison of M_1 which imposes $R\beta = r$ to M_2 which does not have this restriction while the second involves comparing $M_1 : y = X_1\beta_j + \varepsilon_1$ to $M_2 : y = X_2\beta_j + \varepsilon_2$, where X_1 and X_2 contain different explanatory variables, $j = 1, 2$ (See Koop 2003).

Bayes factor is kind of method that is useful in the incorporating of external information into evidence about a model. Thus, Bayes factor through a SDDR will be derived for Bayesian to aid comparison with the frequentist approach. Therefore, this work hereby provides a comparative analysis on the performance of Bayesian using a Bayes factor and classical way of model comparison using a nested model to know the strength of these methods.

The paper is organized as follows. Section 2 gives the regression model and different Bayesian method of model comparison most especially the Bayes factor using Savage-Dickey Density Ratio (SDDR). It also gives a simple decision-theoretic justification model comparison. Section 3 provides procedures for numerical analysis and how the data will be analyzed. In Section 4, results of both classical and Bayesian are presented with decisions. Section 5 concludes the paper.

2. Materials and Methods

In this section, regression model and model comparison methods in both Classical and Bayesian approaches will be given.

2.1. Regression model

Consider the following regression model

$$y = \mathbf{x}\beta + u, \quad (1)$$

where y is an $n \times 1$ vector of observations, β is a $k \times 1$ vector of unknown regression parameters, \mathbf{x} is an $n \times k$ observed matrix of the regression, and u is an $n \times 1$ vector of random errors which is normally distributed with mean zero and constant variance σ^2 . And, $h = 1/\sigma^2$, where h is the precision.

In econometrics, two kinds of models can be used for model comparison: (a) Nested model (b) Non-nested model. The nested model for model comparison involving equality restrictions using M_1 which has the form

$$M_1 : Q\beta = q,$$

against M_2 which does not impose restrictions given as

$$M_2 : Q\beta \neq q,$$

will be used in this study, where Q is $P \times k$ matrix and q is P -vector.

Equation (1) allows for any R linear equality restrictions on the regression coefficient β and it is also assumed that $\text{rank}(Q) = R$. Hence, the two models as suggested by Raftery and Lewis (1996) can be simply defined as

$$H_0 : M_1 : \beta_i = 0 \text{ (reduced model)}, \quad (2)$$

$$H_1 : M_2 : \beta_i \neq 0 \text{ (unreduced model), } i = 0, 1, \dots \quad (3)$$

2.2. Classical model comparison method

Assume the regression model given in (1) is used and also follow the two models given in (2) and (3). Hence, the decision rule for testing the nested model of classical approach is stated below.

Decision rule: if $P < 0.05$, reject H_0 and that means M_2 is supported than M_1 for data under consideration.

2.3. Bayesian model comparison method

There are so many methods that can be use for model comparison in Bayesian context, but some popular methods are however given below:

- (i) Deviance Information Criterion (DIC),
- (ii) Penalized expected deviance,
- (iii) Bayes factor.

Deviance Information Criterion (DIC): It is a hierarchical generalization of AIC for models with weak prior information proposed by Spiegelhalter et al. (2002). This method is an approximation to penalized loss function based on deviance. However, this method was criticized due to under-penalizing of complex models especially in disease mapping situation (See Plummer 2008). The method is simply defined as

$$DIC = \bar{D} + p_D = D(\bar{\theta}) + 2p_D,$$

where \bar{D} is the measure of fit and p_D is the measure of complexity.

Penalized expected deviance: It is Bayesian model comparison method proposed by Plummer (2008). This method was proposed to overcome the problem of when the number of parameters is smaller than number of regressors in a regression model. This estimator is given as

$$L^e(Y, Z) = -2 \int \log\{p(Y | \theta)\} P(\theta | Z) d\theta,$$

where L^e is the loss function for expected deviance.

Bayes Factor: This approach of model comparison is known to be formal and widely accepted. It involves comparing the Posterior odds (PO) ratio of models. The PO is used for comparing two models.

Given two models, say 1 and 2, the PO can be simply stated as:

$$PO_{12} = \frac{P(M_1 | y)}{P(M_2 | y)} = \frac{P(y | M_1)P(M_1)}{P(y | M_2)P(M_2)} = BF_{12} = \frac{P(M_1)}{P(M_2)} = \frac{P(Q\beta = q | y)}{P(Q\beta \neq q | y)},$$

where $P(M_1)$ and $P(M_2)$ are prior model probabilities for models M_1 and M_2 , respectively. $P(y | M_1)$ and $P(y | M_2)$ are marginal likelihood for models M_1 and M_2 , respectively. $P(M_1 | y)$ and $P(M_2 | y)$ are posterior model probabilities for models M_1 and M_2 , respectively.

The Bayes factor, BF_{12} quantifies the weight of evidence in favor of null hypothesis.

2.4. Savage-Dickey Density Ratio (SDDR)

Savage-Dickey Density Ratio (SDDR) as proposed by Dickey (1971) is a generalization of Bayes factor used to compare nested models. It is potentially applicable in variety of applications.

If the priors in two models M_1 and M_2 satisfy the condition below:

$$P(\gamma | \tau = \tau_0, M_2) = P(\tau | M_1),$$

then BF_{12} , the Bayes factor for comparing M_1 to M_2 , will has the form:

$$BF_{12} = \frac{P(\tau = \tau_0 | y, M_2)}{P(\tau = \tau_0 | M_2)},$$

where $P(\tau = \tau_0 | y, M_2)$ and $P(\tau = \tau_0 | M_2)$ are the unrestricted posterior and prior for τ evaluated at the point τ_0 (See Koop 2003).

One major advantage of Savage-Dickey density ratio is that, it involves only M_2 and there is no need to worry about developing methods for posterior inference using M_1 .

Thus, the Savage-Dickey Density Ratio (SDDR) as proposed by Dickey (1971), Verdinelli and Wasserman (1995) for this work will take the form

$$BF_{12} = \frac{P(\beta = \beta_0, h | y, M_2)}{P(\beta = \beta_0, h | M_2)}.$$

For a Bayes Factor (BF) where M_1 imposes equality restriction that is, $\beta_i = 0$, we have

$$BF_{12} = \frac{P(\beta_i, h | y, M_2)}{P(\beta_i, h | M_2)} = \frac{P(\beta_i = 0, h | y, M_2)}{P(\beta_i = 0, h | M_2)},$$

where

$$P(\beta_i, h | y, M_2) = \frac{P(\beta_i, h | M_2)L(\beta_i, h)}{P(y | M_2)} = \frac{P(\beta_i, h | M_2)L(\beta_i, h)}{\iint \pi(\beta_i, h)L(\beta_i, h)d\beta_i dh}. \quad (4)$$

Integrating (4) with respect to h gives:

$$P(\beta_i | y, M_2) = \int P(\beta_i, h | y, M_2)dh = \int \left[\frac{\pi(h | \beta_i)\pi(\beta_i)L(\beta_i, h)}{\iint (\beta_i, h)L(\beta_i, h)d\beta_i dh} \frac{\pi(h | \beta_i)\pi(\beta_i)L(\beta_i, h)}{\iint (\beta_i, h)L(\beta_i, h)d\beta_i dh} \right] dh, \quad (5)$$

Divide (5) by $P(\beta_i | M_2)$ and evaluate both at $\beta_i = 0$ implies:

$$\begin{aligned} \frac{P(\beta_i = 0 | y, M_2)}{P(\beta_i = 0 | M_2)} &= \int \left[\frac{\pi(h | \beta_i = 0)\pi(\beta_i = 0)L(h, \beta_i = 0)}{\pi(\beta_i = 0)\iint \pi(\beta_i, h)L(\beta_i, h)d\beta_i dh} \right] dh = \int \left[\frac{P(h | M_1)L(h, \beta_i = 0)}{\iint \pi(\beta_i, h)L(\beta_i, h)d\beta_i dh} \right] dh, \\ &= \frac{P(y | M_1)}{P(y | M_2)} = BF_{12} = K. \end{aligned}$$

Decision rule: According to Jeffreys (1946) and Kass and Raftery (1995), if $K < 1$, reject H_0 and that means M_2 is supported by the data under consideration than M_1 .

3. Numerical Analysis

3.1. Simulation study

The model for this study is given as:

$$y = 0.1 + 0.4x_1 + 2x_2 + 4.5x_3 + u. \quad (6)$$

The following will be used for simulation of data in the study:

(i) The explanatory variables will be generated as: $x_i \sim N(0, 1)$, $i = 1, 2, 3$.

(ii) The error components is obtained as: $u \sim N(0, h^{-1})$, where $h = 1$.

(iii) The explanatory variables, error term and regression coefficients specified in (6) will be used to generate the dependent variable.

- (iv) The sample sizes are $N=10, 30, 100, 500, 700$ and $1,000$.
- (v) The sample sizes will be replicated $100,000$.

3.2. Real data analysis

Here, the frequentist and Bayesian approaches for model comparison in a nested regression model are applied to a real data. Data on US defence budget outlays for 20 years from 1962 to 1981 are used to know the effects on Gross National Product (GNP), US military sales/assistance and aerospace industry sales on it. The defence budget-outlay, GNP, US military sales/assistance, and aerospace industry sales are measured in \$ billions. The data is sourced from Gujarati and Porter (2005) through various government publications of United States.

Hence, we consider the regression model given as

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3 + u,$$

where y is the defence budget-outlay, x_1 is the GNP, x_2 is the US military sales/assistance and x_3 is the aerospace industry sales. For the simulation study and real life data, the model to be considered for comparison is given as

$$M_1 : \beta_i = 0 \text{ (reduced model),}$$

$$M_2 : \beta_i \neq 0 \text{ (unreduced model), } i = 1, 2, 3.$$

4. Results and Discussion

The objective of this study is to compare frequentist and Bayesian approaches for model comparison in a nested regression model. This section gives the results obtained from Bayesian and classical approaches with various decisions for parameters. Here, M_2 is an unreduced model and represents a good model. The Bayes factor, (K) gives the result for Bayesian using Savage Dickey Density Ratio (SDDR) while the p-value for the t-distribution is also presented for frequentist approach.

Table 1 Model comparison of nested model for Bayesian and classical methods with sample size of 10

Parameter	Estimators	Estimates ($K/\Pr(t)$)	Decision
β_0	Bayesian	0.000859	M_2
	Classical	0.151000	M_1
β_1	Bayesian	0.001174	M_2
	Classical	0.067300	M_1
β_2	Bayesian	0.000801	M_2
	Classical	1.12E-05	M_2
β_3	Bayesian	0.000123	M_2
	Classical	1.12E-05	M_2

Table 2 Model comparison of nested model for Bayesian and Classical methods with sample size of 30

Parameter	Estimators	Estimates ($K/\Pr(t)$)	Decision
β_0	Bayesian	2.4E-14	M_2
	Classical	0.429700	M_1
β_1	Bayesian	3.31E-14	M_2
	Classical	0.574500	M_1
β_2	Bayesian	2.18E-14	M_2
	Classical	0.043400	M_2
β_3	Bayesian	1.12E-14	M_2
	Classical	1.67E-07	M_2

Results obtained from both Tables 1 and 2 show that Bayesian method of model comparison picked a good model for all the parameters while the classical method only picked a good model for parameters β_2 and β_3 for sample size of 10 and 30.

Table 3 Model Comparison of nested model for Bayesian and classical methods with sample size of 100

Parameter	Estimators	Estimates ($K/\Pr(t)$)	Decision
β_0	Bayesian	0.00E+00	M_2
	Classical	0.783300	M_1
β_1	Bayesian	0.00E+00	M_2
	Classical	0.001900	M_2
β_2	Bayesian	0.00E+00	M_2
	Classical	2.00E-16	M_2
β_3	Bayesian	0.00E+00	M_2
	Classical	0.00E+00	M_2

In Tables 3-6, Bayesian method of model comparison give more evidence against the reduced model M_1 for all the parameters considered. The classical approach gives evidence of support for most especially parameter β_2 and β_3 in Tables 3-6. It was observed that parameter β_0 does not give support for the unreduced model for classical method of model comparison for all the sample sizes as shown in Tables 1-6.

Table 4 Model Comparison of nested model for Bayesian and classical methods with sample size of 500

Parameter	Estimators	Estimates ($K/\Pr(t)$)	Decision
β_0	Bayesian	4.04E-67	M_2
	Classical	0.866080	M_1
β_1	Bayesian	5.76E-68	M_2
	Classical	0.003390	M_2
β_2	Bayesian	3.31E-67	M_2
	Classical	4.34E-05	M_2
β_3	Bayesian	8.01E-68	M_2
	Classical	<2.05	M_2

Table 5 Model Comparison of nested model for Bayesian and classical methods with sample size of 700

Parameter	Estimators	Estimates ($K/\Pr(t)$)	Decision
β_0	Bayesian	0	M_2
	Classical	0.301000	M_1
β_1	Bayesian	0	M_2
	Classical	0.247000	M_1
β_2	Bayesian	0	M_2
	Classical	2.00E-16	M_2
β_3	Bayesian	0	M_2
	Classical	2.00E-16	M_2

Table 6 Model Comparison of nested model for Bayesian and classical methods with sample size of 1,000

Parameter	Estimators	Estimates ($K/\Pr(t)$)	Decision
β_0	Bayesian	0	M_2
	Classical	0.453120	M_1
β_1	Bayesian	0	M_2
	Classical	0.005230	M_1
β_2	Bayesian	0	M_2
	Classical	2.00E-16	M_2
β_3	Bayesian	0	M_2
	Classical	2.00E-16	M_2

Table 7 Model Comparison of nested model for Bayesian and classical methods in real life data application

Parameter	Estimators	Estimates ($K/\Pr(t)$)	Decision
β_0	Bayesian	0.0000	M_2
	Classical	0.0000	M_2
β_1	Bayesian	0.0008	M_2
	Classical	0.0300	M_1
β_2	Bayesian	0.0000	M_2
	Classical	0.1450	M_1
β_3	Bayesian	0.0000	M_2
	Classical	0.0000	M_2

In results obtained from the real life data application, Bayesian method of model comparison give more evidence against the reduced model M_1 for all the parameters considered. This also agreed with the results obtained in simulation.

5. Conclusion

Overtime, different work had been carried out by facilitating comparison between the two modelling approaches, that is, frequentist and Bayesian methods in the area of parameter estimation. It is of interest for model builders to know the strength of the two methods in the area of model comparison.

In this paper, comparison between the Bayesian and frequentist approaches using a nested model was carried out to know the strength of the two methods. Bayes factor through a Savage-Dickey density ratio when parameters are evaluated at zero was provided for Bayesian method in order to facilitate comparison.

Data were simulated for different sample sizes and each of the sample sizes was replicated while a real data set was also applied. For all the sample sizes considered, Bayesian performs well than the frequentist approach by giving an overwhelming evidence support for model M_2 for all the parameters.

Therefore, the Bayesian method of model comparison outperformed the frequentist method for nested model. Thus, Bayesian method using a Savage-Dickey density ratio is preferred to frequentist approach where a researcher has choice between the two approaches.

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